

Efficient Simulation for Rare Events

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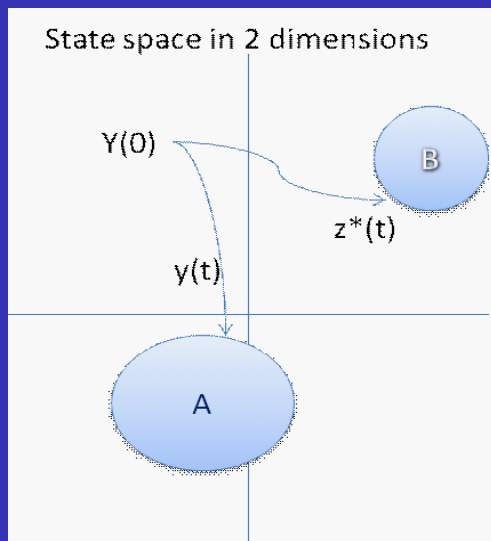
- **Review**
- Stylized Example: Light-tailed Random Walks
- Control Theory, Harmonic Functions and Doob's h -transform
- Lyapunov Inequalities
- Stylized Example: Heavy-tailed Random Walks

Generic Rare-event Simulation Problem

- A generic rare-event estimation problem:

$$P(\text{Hit } B \text{ prior to } A)$$

Generic Rare-event Simulation Problem



- Under suitable light-tailed assumptions:

$$\begin{aligned} & \Delta \log P (\text{Hit } B \text{ prior to } A) \\ \approx & - \inf \{ J(z) : z(\cdot) \text{ is path that hits } B \text{ prior to } A \} \end{aligned}$$

Description of State-independent Approach: Review

- Under suitable light-tailed assumptions:

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- $J(\cdot)$ is the action function and $z^*(\cdot)$ is the “optimal path”

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- $J(\cdot)$ is the action function and $z^*(\cdot)$ is the “optimal path”
- Tracking optimal path apply exponential tilting at time t to follow $\dot{z}^*(t)$

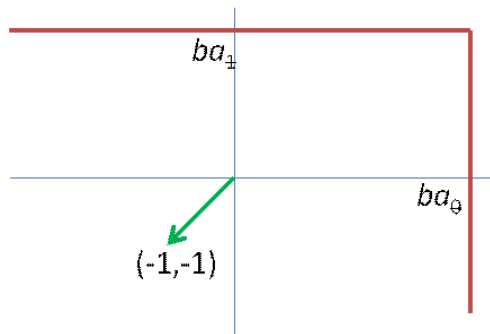
Counter-examples

- **Model:** $Y_t = (-1, -1)t + X_t$; X_t is Brownian motion &
 $B_b = \{(x, y) : x \geq a_0 b \text{ or } y \geq a_1 b\}$

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- **Estimate:** $u(b) = P(T_{B_b} < \infty)$ as $b \nearrow \infty$

First Passage Time Problem in two dimensions



- **Consider estimating:** $P(T_b < \infty)$, $T_b = \inf\{n \geq 0 : S_n > b\}$

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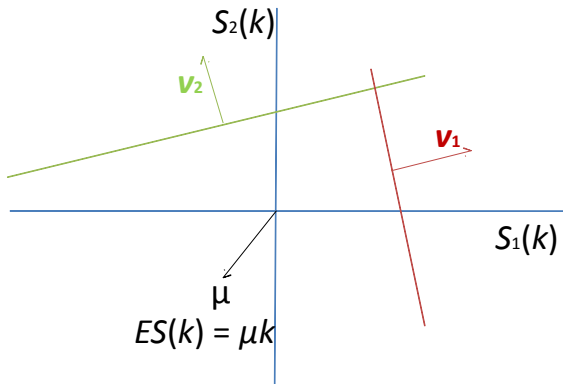
- **No clear way to mimic zero variance change-of-measure!**
- **No clear way to apply the systematic approach!**

State – dependent importance sampling

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Back to Counter-example in Light-tailed Setting

- Two dimensional random walk
- $A_n = \{s : v_2^T s \geq 1\}$ and $B_n = \{s : v_1^T s \geq 1\}$



- Efficiently estimate as $n \nearrow \infty$

$$u_n(0) = P_0[S_k/n \text{ hits } A \text{ OR } B \text{ Eventually}]$$

Elements of Associated Large Deviations

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- Note $EZ_k^{(1)} = v_1^T \mu < 0$ and $EZ_k^{(2)} = v_2^T \mu < 0$
- Assume there are $\theta_1^*, \theta_2^* > 0$ such that

$$\begin{aligned} E \exp(\theta_1^* Z_k^{(1)}) &= 1 \quad \& \quad E \exp(\theta_2^* Z_k^{(2)}) = 1 \\ E[\exp(\theta_1^* Z_k^{(1)}) Z_k^{(1)}] &< \infty \quad \& \quad E[\exp(\theta_2^* Z_k^{(2)}) Z_k^{(2)}] < \infty \end{aligned}$$

Large Deviations for the Stylized Example

- As $n \nearrow \infty$

$$\begin{aligned}u_n(x) &= P_x[W_n(t) \text{ hits } A \text{ OR } B] \\ &\sim c_1 \exp(-n\theta_1^*(1 - v_1^T x)) + c_2 \exp(-n\theta_2^*(1 - v_2^T x)) \\ &= \exp(-nh(x) + o(n)),\end{aligned}$$

where

$$h(x) = \min[\theta_1^*(1 - v_1^T x), \theta_2^*(1 - v_2^T x)].$$

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- Game interpretation \rightarrow explains Isaacs equation name...

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- **Solution:** $\lambda = -\partial g(w)/2$ and $\psi(-\partial g(w)/2) = 0$ subject to $g(w) = I(w \in A \cup B)$

Harmonic Functions

- $u_n(w) = 1$ on $A \cup B$ & harmonic:

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$$P^*(X_{k+1} \in dy | S_k = nw) = f(y) \frac{u_n(w + X/n)}{u_n(w)}$$

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- Constant likelihood ratio \implies zero variance

$$u_n(S_0/n) = \frac{u_n(S_0/n)}{u_n(S_1/n)} \times \frac{u_n(S_1/n)}{u_n(S_2/n)} \times \dots \times \frac{u_n(S_{T_{A \cup B}-1}/n)}{u_n(S_{T_{A \cup B}}/n)}$$

- Zero-variance sampler

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$$\begin{aligned} & \tilde{P} (Y_{k+1} \in dy | S_k = nw) \\ & \approx f(y) \exp(-n[h(w + y/n) - h(w)]) \\ & \approx f(y) \exp(-\partial h(w) \cdot y) \end{aligned}$$

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- **Equivalent to Isaacs equation with $g(w) = 2h(w) \rightarrow$ best asymptotic rate**

Dynamic Programming, Isaacs Equation, Harmonic Functions: Summary

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- Consider sampler

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- Second moment of estimator

$$s(w) = E[r(w, w + X/n)s(w + X/n)]$$

subject to $s(w) = 1$ for $w \in A \cup B$.

Lemma

B. & Glynn '08: Lyapunov inequality

$$v(w) \geq E[r(w, w + X/n)v(w + X/n)]$$

subject to $v(w) \geq 1$ for $w \in A \cup B$. Then, $v(w) \geq s(w)$.

- **How to use the result?** 1) Identify a change-of-measure, 2) use heuristic / approx. to force $v(w) \approx u_n(w)^2$.

Lyapunov Inequalities and Subsolutions

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- Yields subsolution to the Isaacs equation (note smoothness)

$$\psi(-\partial g(w)/2) \leq 0 \text{ s.t. } g(w) \leq 0, w \in A \cup B$$

The Mollification

- Back to random walk example

$$\begin{aligned}h(w) &= \min[\theta_1^*(1 - v_1^T w), \theta_2^*(1 - v_2^T w)] \\ &= -\max[\theta_1^*(v_1^T w - 1), \theta_2^*(v_2^T w - 1)]\end{aligned}$$

NOT smooth...

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- Mollification:

$$\begin{aligned}h_\varepsilon(w) \\ = -\varepsilon \log[\exp(\theta_1^*(v_1^T w - 1)/\varepsilon) + \exp(\theta_2^*(v_2^T w - 1)/\varepsilon)]\end{aligned}$$

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- Implementation via mixtures:

$$\begin{aligned}-\partial h_\varepsilon(w) &= \theta_1^* v_1^T \frac{\eta_1^\varepsilon(w)}{\eta_1^\varepsilon(w) + \eta_2^\varepsilon(w)} + \theta_2^* v_2^T \frac{\eta_2^\varepsilon(w)}{\eta_1^\varepsilon(w) + \eta_2^\varepsilon(w)}, \\ \eta_1^\varepsilon(w) &= \exp(\theta_1^*(v_1^T w - 1)/\varepsilon), \\ \eta_2^\varepsilon(w) &= \exp(\theta_2^*(v_2^T w - 1)/\varepsilon).\end{aligned}$$

The Form of a Typical Optimality Statement...

Theorem (Dupuis & Wang '07)

Let $g_{\varepsilon_n}(w) = 2h_{\varepsilon_n}(w)$ and assume that $n\varepsilon_n \rightarrow \infty$ apply corresponding sampler. Then,

$$2\text{nd Moment of Est.} = \exp(-2nh(w) + o(n)).$$

- Select $\varepsilon_n = 1/n$

$$\eta_1(w) = \exp(n\theta_1^*(v_1^T w - 1))$$

$$\eta_2(w) = \exp(n\theta_2^*(v_2^T w - 1))$$

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- Mixture sampler from density $\tilde{f}(x)$

$$\frac{\tilde{f}(x)}{f(x)} = \frac{\eta_1(w)}{\eta_1(w) + \eta_2(w)} \exp(\theta_1^* v_1^T x) + \frac{\eta_2(w)}{\eta_1(w) + \eta_2(w)} \exp(\theta_2^* v_2^T x)$$

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- Lyapunov function

$$v(w) = (\eta_1(w) + \eta_2(w))^2 \geq 1$$

for $v_1^T w \geq 1$ OR $v_2^T w \geq 1$... BOUNDARY CONDITION OK!

A Lyapunov Inequality



$$\begin{aligned}v(w) &= [\eta_1(w) + \eta_2(w)]^2 \\ \eta_1(w + X/n) &= \eta_1(w) e^{\theta^* v_1^T X} \\ \eta_2(w + X/n) &= \eta_2(w) e^{\theta^* v_2^T X}\end{aligned}$$

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$$\begin{aligned}& E \frac{v(w + X/n)}{v(w)} \frac{1}{\frac{\eta_1(w)}{\eta_1(w) + \eta_2(w)} e^{\theta_1^* v_1^T X} + \frac{\eta_2(w)}{\eta_1(w) + \eta_2(w)} e^{\theta_2^* v_2^T X}} \\ &= E \frac{\eta_1(w) \exp(\theta_1^* v_1^T X) + \eta_2(w) \exp(\theta_2^* v_2^T X)}{\eta_1(w) + \eta_2(w)} = 1.\end{aligned}$$

Theorem (B., Glynn and Leder (2009))

One can take $\varepsilon = 1/n$ as mollification parameter & in fact this is the optimal choice as it gives bounded coef. of variation

$$2nd \text{ Moment} \leq (v_1(0) + v_2(0))^2 \leq cu_n(0)^2.$$

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- Let X_1, X_2, \dots are heavy-tailed (TBD) and $EX_i = \mu < 0$
- $S_n = X_1 + \dots + X_n$ given $S_0 = s$
- Object of interest:

$$u_b(s) = P_s(T_b < \infty) = \frac{\int_{b-s}^{\infty} P(X_i > u) du}{-\mu} (1 + o(1))$$

as $b - s \nearrow \infty$.

- Asymptotics: Pakes, Veraverbeke, Cohen... see text of Asmussen '03

- Rich theory for heavy-tailed random walks based on *subexponentiality*

$$P(X_1 + X_2 > b) = 2P(X_1 > b)(1 + o(1))$$

as $b \rightarrow \infty$.

- *Focus on regularly varying distributions* (basically *power-law* type)

$$P(X_1 > t) = t^{-\alpha} L(t)$$

for $\alpha > 1$ and $L(t^\beta) / L(t) \rightarrow 1$ as $t \nearrow \infty$ for each $\beta > 0$.

- **Importance sampling:** Markov kernel $K(\cdot)$

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- **Recall to get good variance:** *Select an IS that mimics the conditional distribution.*

Description of the Conditional Distribution

Theorem

(Asmussen and Kluppelberg): Conditional on $T_b < \infty$, we have that

$$\left(\frac{S_{uT_b}}{T_b}, \frac{S_{T_b} - b}{b}, \frac{T_b}{b} \right) \Longrightarrow (\mu u, Z_1, Z_2)$$

on $D(0, 1) \times R \times R$ as $b \nearrow \infty$, where Z_1 and Z_2 are Pareto with index $\alpha - 1$.

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- **Interpretation:** Prior to ruin, random walk has drift μ and a large jump of size b occurs suddenly in $O(b)$ time...
- So, given that a jump hasn't occurred by time k , then $S_k \approx \mu k$ and the chance of reaching b in the next increment *given* that we eventually reach b ($T_b < \infty$)

$$\frac{P(X > b - \mu k)}{\int_0^\infty P(X > b - \mu u) du} \approx \frac{-\mu P(X > b - \mu k)}{\int_b^\infty P(X > s) du} = O\left(\frac{1}{b}\right).$$

Good Family of Changes of Measure for Heavy Tails

- **Change-of-measure:** Here s is current position of the walk, f is the density

$$f_{X|s}(x|s) = p(s) \frac{f_X(x) I(x > a(b-s))}{P(X > a(b-s))} + (1-p(s)) \frac{f_X(x) I(x \leq a(b-s))}{P(X > a(b-s))}$$

- In other words, $s_0 = s$ and $s_1 = s_0 + x$

$$r(s_0, s_1)^{-1} = p(s_0) \frac{I(s_1 - s_0 > a(b-s_0))}{P(X > a(b-s_0))} + (1-p(s_1)) \frac{I(s_1 - s_0 \leq a(b-s_0))}{P(X \leq a(b-s_0))}$$

- Introduced by Dupuis, Leder and Wang '06 for finite sums...

- Recall Lyapunov inequality

Lemma (B. & Glynn '08)

Suppose that there is a positive function $g(\cdot)$ such that

$$E_s^K \left(\frac{g(S_1) r(s, S_1)^2}{g(s)} \right) = E_s \left(\frac{g(S_1) r(s, S_1)}{g(s)} \right) \leq 1$$

for all $s \leq b$ and $g(s) \geq 1$ for $s > b$. Then,

$$E_s^K Z^2 = E_s^K \left(\prod_{j=1}^{T_b-1} r(S_j, S_{j+1})^2 I(T_b < \infty) \right) \leq g(s).$$

Constructing Lyapunov Functions for Heavy-tails

- Wish to achieve *strong efficiency*, so we pick (for some $\kappa > 0$)

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- **Must select κ and θ to verify Lyapunov inequality**

- **Testing the Inequality on $g(s) < 1$ (note that $g \leq 1$):**

$$\begin{aligned} & E_s \left(\frac{g(S_1) r(s, S_1)}{g(s)} \right) \\ = & \frac{E(g(s+X); X > a(b-s)) P(X > a(b-s))}{p(s) g(s)} \\ & + \frac{E(g(s+X); X \leq a(b-s)) P(X \leq a(b-s))}{(1-p(s)) g(s)} \end{aligned}$$

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Testing the Lyapunov Inequality

- **Testing the Inequality on $g(s) < 1$ (note that $g \leq 1$):**

$$\begin{aligned} & E_s \left(\frac{g(S_1) r(s, S_1)}{g(s)} \right) \\ & \leq \frac{P(X > a(b-s))^2}{p(s)g(s)} + \frac{E(g(s+X); X \leq a(b-s))}{(1-p(s))g(s)} \\ & \approx \frac{a^{-\alpha} P(X > a(b-s))}{\theta\kappa \int_{b-s}^{\infty} P(X > u) du} + 1 + 2(\mu + \theta) \frac{P(X > (b-s))}{\left(\int_{b-s}^{\infty} P(X > u) du \right)} \end{aligned}$$

- **NOTE:** crucial that $\mu < 0$! Pick θ small and κ large

$$\frac{a^{-\alpha}}{\theta\kappa} + 2\theta + 2\mu \leq 0$$

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- Construct Lyapunov function from course large deviations analysis
- Lyapunov inequalities guide selection of mollification parameters and guarantee good performance

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