
On the Initial Transient Problem for Steady-State Simulation

Peter W. Glynn
Stanford University

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Outline of Talk

1. The **Steady-State Simulation** Problem
 - What is it?
 - Why is it relevant?
 - Challenges
2. The **Initial Transient** Problem
 - Is it a serious problem?
 - When is it a serious problem?
3. The Basic Approaches to the Initial Transient Problem
 - Simulating a stationary version
 - Identifying the initial transient period
 - Low Bias estimators

Steady-State Simulation

$Y(t)$ = rate at which “reward”
increases at time t

Assume there exists a (deterministic) constant α
such that

$$\frac{1}{t} \int_0^t Y(s) ds \Rightarrow \alpha$$

as $t \rightarrow \infty$.

Goal: Compute α

Examples

$Y(t)$ = **number-in-system (in queue) at time t**

$Y(t)$ = **work-in-process (in manufacturing system)
at time t**

$Y(t) = I(Q(t) \geq b)$

Why the emphasis on steady-state simulation?

To compute $\alpha(t, x) = \mathbb{E}_x f(X(t))$:

$$\alpha'(t) = A\alpha(t)$$

$$\alpha(0) = f$$

To compute $\alpha = \sum_x \pi(x) f(x)$:

$$0 = \pi A$$

$$\mathbf{s/t} \sum_x \pi(x) = 1$$

$$\alpha = \pi f$$

Analytical and computational tractability

Modeling Issues

(favoring use of a steady-state formulation)

- ▶ no need to specify time horizon
- ▶ no need to specify initial distribution
- ▶ appropriate in many applications

Steady-State Simulation: Challenge 1

- ▶ The steady state mean α involves the “infinite time behavior” of $(Y(s) : s \geq 0)$
- ▶ How do we compute α based on a finite-horizon simulation?
- ▶ Mathematically: $E \frac{1}{t} \int_0^t Y(s) ds \neq \alpha$
- ▶ Conceptually: Initial condition is atypical of steady-state
 ↓
 “initial transient”
 ↓
 “initial bias”

Steady-State Simulation: Challenge 2

In view of Challenge 1, choose time horizon t **large**



One observation of process Y over $[0, t]$



How to compute an estimator for the **variance** of

$$\frac{1}{t} \int_0^t Y(s) ds ?$$

A Connection between the Two Challenges

$$E_x f(X(t)) = \alpha + a(x)e^{-\eta t}(1 + o(1))$$

and

$$\begin{aligned} \text{cov}(f(X^*(0)), f(X^*(t))) &= E_{\pi} f_c(X(0)) f_c(X(t)) \\ &= E_{\pi} f_c(X(0)) \cdot E[f_c(X(t)) | X(0)] \\ &= E_{\pi} f_c(X(0)) \cdot a(X(0)) e^{-\eta t} (1 + o(1)) \end{aligned}$$

- ▶ used intensively
- ▶ doesn't work in presence of non-Markov processes

Today: We focus on **Challenge 1**

“The Initial Transient Problem”

How Serious a problem is the Initial Transient?

For **geometrically ergodic** (Markov) processes,

$Y(t) \Rightarrow Y(\infty)$ as $t \rightarrow \infty$ and

$$\mathbf{E}Y(t) = \mathbf{E}Y(\infty) + O(e^{-\eta t})$$

so

$$\begin{aligned} \mathbf{E} \frac{1}{t} \int_0^t Y(s) ds - \alpha &= \frac{1}{t} \int_0^\infty \{\mathbf{E}Y(s) - \mathbf{E}Y(\infty)\} ds - \frac{1}{t} \int_t^\infty \{\mathbf{E}Y(s) - \mathbf{E}Y(\infty)\} ds \\ &= b/t + O(e^{-\eta t}) \end{aligned}$$

Also, $t^{1/2} \left(\frac{1}{t} \int_0^t Y(s) ds - \alpha \right) \Rightarrow \sigma N(0, 1)$

$$\text{var} \left(\frac{1}{t} \int_0^t Y(s) ds \right) \sim \sigma^2/t$$

Conclusion

- ▶ Initial bias effect is **small** relative to **sampling variability**

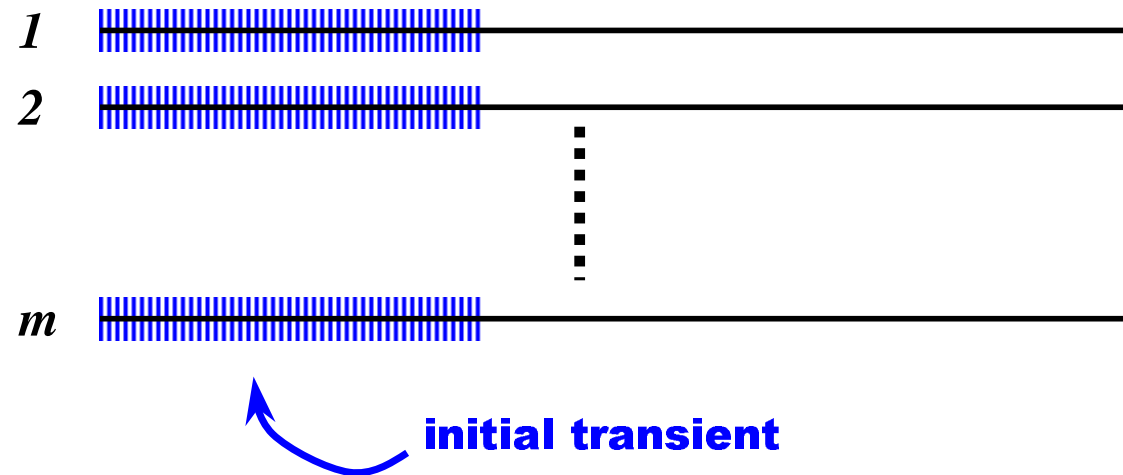
- ▶ For most OR-related steady-state simulations, initial transient is not serious and **can be ignored**

Why: Any time-horizon large enough to make the sampling variability small effectively eliminates the effect of the initial transient.

There are exceptions

- ▶ **parallel simulation**
- ▶ **long-range dependent processes**
- ▶ **multi-modal behavior / nearly decomposable systems**
- ▶ **high dimensional systems**

Parallel Simulation



► **MSE analysis:**

$$\text{variance} \left(\frac{1}{m} \sum_{i=1}^m \frac{1}{t} \int_0^t Y_i(s) ds \right) \sim \frac{\sigma^2}{mt}$$

$$\text{bias} \left(\frac{1}{m} \sum_{i=1}^m \frac{1}{t} \int_0^t Y_i(s) ds \right) \sim \frac{b}{t}$$

► **If $m \gg t$, initial transient becomes dominant effect.**

GH 90's

Long-range dependence

- ▶ In the presence of long-range dependence,

$$\text{var} \left(\frac{1}{t} \int_0^t Y(s) ds \right) \sim \frac{\sigma^2}{t^{2-2H}} \quad (1/2 < H < 1)$$

- ▶ How to allocate c units of computer time in single processor setting?

m independent **replications** of length c/m

$$\text{var} \left(\frac{1}{m} \sum_{i=1}^m \frac{1}{c/m} \int_0^{c/m} Y_i(s) ds \right) \approx \frac{\sigma^2 m^{1-2H}}{c^{2-2H}}$$

- ▶ Choose m large ...
- ▶ Same initial bias problem as in **parallel simulation** setting

High-Dimensional Systems

$$\vec{Y}(t) = (Y_1(t), \dots, Y_d(t))$$

independent

$$\|P(Y_i(t) \in \cdot) - P(Y_i(\infty) \in \cdot)\|_y \sim ce^{-\lambda t}$$

$$\|P(\vec{Y}(t) \in \cdot) - P(\vec{Y}(\infty) \in \cdot)\| \sim dce^{-\lambda t} \quad \text{G05}$$

Slows down
convergence rate...

**There are models for which
initial transient is significant**

3 Approaches to Dealing with the Initial Transient

- ▶ **Approach 1** Simulate a stationary version of Y
exact simulation / perfect simulation

- ▶ **Approach 2** Identify the initial transient interval
 $[0, t]$

- ▶ **Approach 3** Modify the estimator so as to reduce
the “initial bias” effect

Approach 1: Simulate a stationary version of Y

$$Y(t) = f(X(t))$$

- ▶ Choose $X(0) \sim \pi$
- ▶ Don't know π !
- ▶ Can we sample from π based on the ability to do a dynamic simulation of X ?
- ▶ Exact/ perfect simulation

$$\begin{aligned}\pi(\cdot) &= \frac{\mathbb{E} \int_0^\tau I(X(s) \in \cdot) ds}{\mathbb{E} \tau} \\ &= \int_0^\infty \mathbb{P}(X(s) \in \cdot \mid \tau > s) \frac{\mathbb{P}(\tau > s) ds}{\mathbb{E} \tau} \\ &= \mathbb{P}(X(Z) \in \cdot) \quad \text{AGT 92}\end{aligned}$$

$$Z \sim \frac{\mathbb{P}(\tau > \cdot)}{\mathbb{E} \tau}$$

-
- ▶ Are there any interesting problem classes for which

$$P(\tau > \cdot) / E \tau$$

is known, but simulating from $\pi(\cdot)$ directly is hard?

- ▶ Yes

$$P^m(x, y) \geq \varepsilon \phi(y)$$

$$P^m(x, y) = \varepsilon \phi(y) + (1 - \varepsilon)q(x, y)$$

$$P(\tau = mk) = \varepsilon(1 - \varepsilon)^{k-1}$$

$\frac{\tau}{m}$ **geometric**

AGT (1992)

Propp/Wilson (1996)

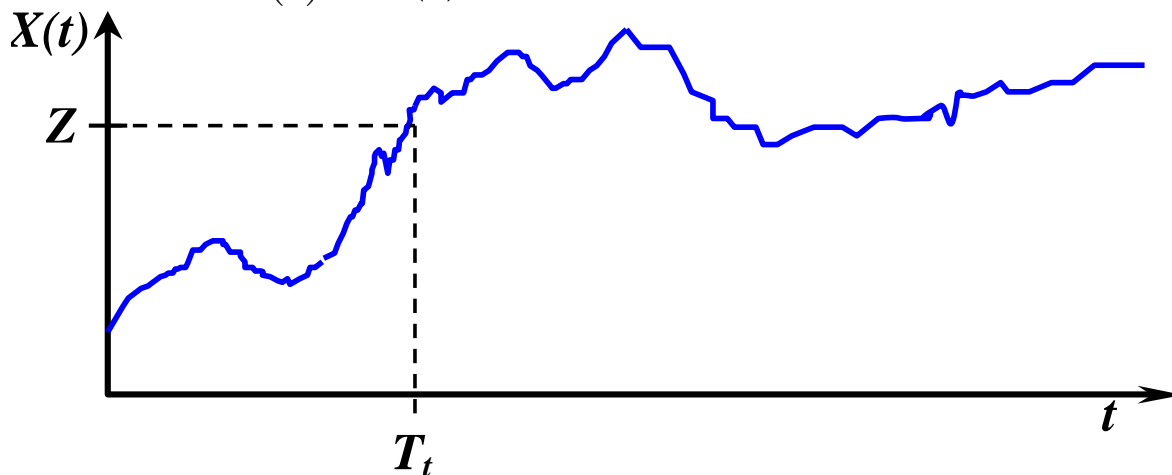
⋮

A Statistically-based alternative

- ▶ Simulate Markov process X over $[0, t]$
- ▶ Compute $\hat{\pi}_t(x) = \frac{1}{t} \int_0^t I(X(s) = x) ds$
- ▶ Generate $Z_t \sim \hat{\pi}_t$ ▶ Put $T_t = \inf\{s \geq 0 : X(s) = Z_t\}$
- ▶ Then,

$$\|(X(s) : T_t \leq s \leq T_t + t) - (X^*(s) : 0 \leq s \leq t)\| \rightarrow 0 \quad \text{G93}$$

- ▶ $\mathbf{E} \left(\frac{1}{t - T(t)} \int_{T(t)}^t f(X(s)) ds \right) = \alpha + O\left(\frac{1}{t^2}\right) \quad (\text{A} + \text{G} 06)$



Approach 2: Identify initial transient interval $[0, t_0]$

- Use **analytic bounds** to compute t_0 for which

$$\|P(Y(t_0) \in \cdot) - P(Y(\infty) \in \cdot)\| < \varepsilon$$

For many **geometrically ergodic** processes,

$$E_x f(X(t)) = \alpha + a(x)e^{-\gamma t + i\theta t}(1 + o(1))$$

where $\gamma > 0$. Here, $z = -\gamma + i\theta$ is second largest (in modulus) eigenvalue of A .

- For highly structured models, one can sometimes compute asymptotics for z (Diaconis)
- For **reversible** systems, one can compute analytic bounds on z
- For **regenerative** systems, one can obtain bounds on

$$\|P(Y(t) \in \cdot) - P(Y(\infty) \in \cdot)\|$$

-
- ▶ Use an approximating analytic model to identify t_0 :

$$Y(t) \stackrel{\mathcal{D}}{\approx} \tilde{Y}(t)$$

Single-server FIFO queue in “heavy-traffic”:

$$Y(t) \stackrel{\mathcal{D}}{\approx} \text{RBM}(t)$$
$$t_0 = c(1 - \rho)^{-2}$$

For infinite-server queue in “heavy-traffic”:

$$Y(t) \stackrel{\mathcal{D}}{\approx} \text{Gaussian process}$$
$$t_0 \approx F^{-1}(1 - \varepsilon)$$

Problems

- ▶ good approximations hard to compute for most models (networks)

Possible remedy:

- simulate simplified model first; determine t_0 based on simulation
 - then simulate real model
- ▶ note that time to stationarity t_0 depends on functional being computed

t_0 for “number-in-system”

$\neq t_0$ for “buffer loss”

γ for “number-in-system”

$\neq \tilde{\gamma}$ for exponential moments of “number-in-system”

Statistically Based Methods

- ▶ **Many proposals have been made**
e.g. discard initial observations until the first one left is neither the maximum nor the minimum of the remaining observations (Conway 1963)
- ▶ **Problematic both in theory and in practice**

A Basic Difficulty

- ▶ In general, one has only one realization of the initial transient on which to base one's statistical analysis.
- ▶ Any rigorous rule needs to “model” the initial transient
e.g., autoregressive process

$$X_{n+1} = \rho X_n + \varepsilon_{n+1}$$

Regenerative Processes

$$a(t) = \mathbf{E} Y(t)$$

$$a = b + F * a \quad (\text{renewal eqn})$$

Solution is

$$a = U * b$$

If b, F are **known**, choose t_0 such that $|a(t_0) - a(\infty)| < \varepsilon$

If b, F are **unknown**, estimate these from cycles simulated:

$$\widehat{F}_n(t) = \frac{1}{n} \sum_{i=1}^n I(\tau_i \leq t)$$

$$\widehat{b}_n(t) = \frac{1}{n} \sum_{i=1}^n b_i(t)$$

Choose \widehat{t} so that

$$|(\widehat{U} + \widehat{b})(\widehat{t}) - (\widehat{U} + \widehat{b})(\infty)| < \varepsilon \quad \text{BG06, KG06}$$

Approach 3: Low Bias Estimators

Attempt 1

$$\mathbb{E} \frac{1}{t} \int_0^t Y(s) ds = \alpha + \frac{b}{t} + O(e^{-\lambda t})$$

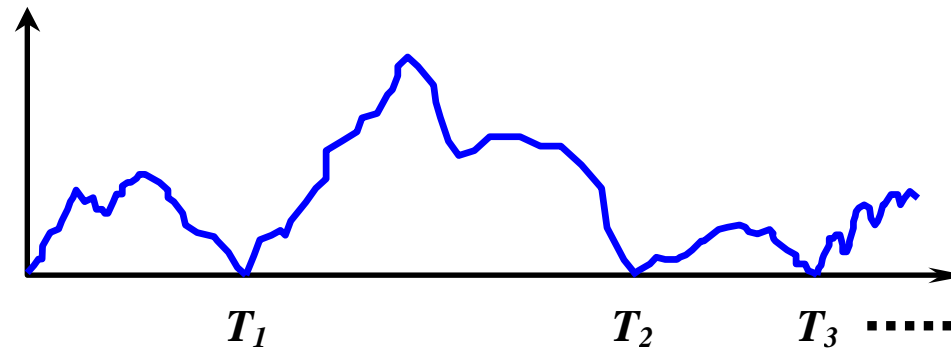
Estimate α , b via a linear regression

Use $\hat{\alpha}$ from linear regression

Doesn't work ...

As for Approach 2, must “model” initial transient

Regenerative Processes



$$Y = (Y(s) : 0 \leq s \leq T_n)$$
$$\left(\int_{T_{i-1}}^{T_i} Y(s) ds, T_i - T_{i-1} : 1 \leq i \leq n \right)$$

iid

- ▶ No initial transient on time scale of regenerative cycles
- ▶ Structure of initial transient highly dependent on time scale used

How does initial bias manifest itself on time scale of regenerative cycles?

$$\alpha = \frac{\mathbb{E} \int_0^{T_1} Y(s) ds}{\mathbb{E} T_1} \triangleq \frac{\mathbb{E} W}{\mathbb{E} T}$$

$$\alpha_n = \frac{\bar{W}_n}{\bar{T}_n} = g(\bar{W}_n, \bar{T}_n)$$

$$\mathbb{E} \alpha_n \neq g(\mathbb{E} \bar{W}_n, \mathbb{E} \bar{T}_n)$$

← non-linearity bias

Solutions:

- ▶ Taylor expand ...

estimate



$$\mathbb{E} \alpha_n = \alpha - \frac{1}{n} \frac{\mathbb{E}[(W - \alpha T)T]}{(\mathbb{E} T)^2} + O(n^{-2}) \quad \text{CL75}$$

- ▶ Jack-knife

On time scale of simulated time:

$$\mathbb{E} \frac{1}{t} \int_0^t Y(s) ds = \alpha + \frac{b}{t} + O(e^{-\lambda t})$$

where

$$b = \frac{\mathbb{E} \int_0^{T_1} s [Y(s) - \alpha] ds}{\mathbb{E} T_1}$$

estimate from
cycles completed by t 

G 90

Typical Theoretical Analysis

$$E \hat{\alpha}_L(t) = \alpha + o(1/\tau)$$

$$\text{var } \hat{\alpha}_L(t) \sim \text{var } \hat{\alpha}(t) \quad \text{as } t \rightarrow \infty.$$

Hence, $\hat{\alpha}_L(t)$ is better...

A More Careful Analysis

$$E \hat{\alpha}(t) = \alpha + \frac{b}{t} + o(1/t)$$

$$E \hat{\alpha}_L(t) = \alpha + o(1/t)$$

$$\text{var } \hat{\alpha}(t) = \frac{\sigma^2}{t} + \frac{c_1}{t^2} + o(1/t^2)$$

$$\text{var } \hat{\alpha}_L(t) = \frac{\sigma^2}{t} + \frac{c_2}{t^2} + o(1/t^2)$$

$$\text{MSE}(\hat{\alpha}(t)) = \frac{\sigma^2}{t} + \frac{c_1 + b^2}{t^2} + o(1/t^2)$$

$$\text{MSE}(\hat{\alpha}_L(t)) = \frac{\sigma^2}{t} + \frac{c_2}{t^2} + o(1/t^2)$$

Need to check that $c_2 < c_1 + b^2$ in single replication setting; When a large number of processors are used, $\hat{\alpha}_L(t)$ is clearly better A+G 07

Conclusions

- ▶ **There are problems in which initial transient can be a serious issue**
- ▶ **There are rigorously supported methods for reducing the effect of the initial transient
even in single replication setting!**
- ▶ **Ideas apply even in nonstationary settings, e.g., when to initialize simulations focused on “rush hours”**