

Efficient Simulation for Rare Events

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- **Introduction**
- A Simple Random Walk Example
- Systematic Approach
- Testing Efficiency
- Counter-examples, Heavy-tails and Beyond

Rare Event Analysis: Areas of Applicability

Rare events are consequential in many areas

- Insurance / Finance

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- Reliability models

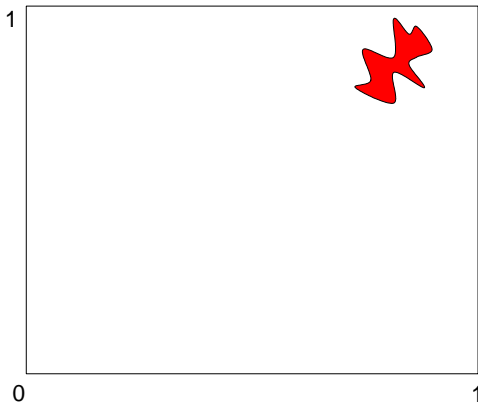
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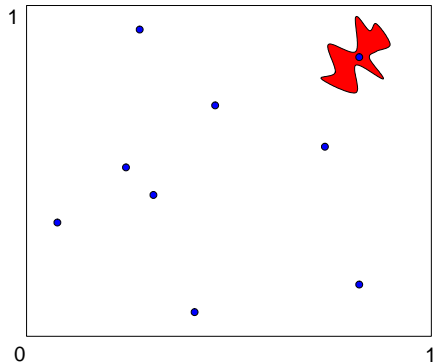
- Insurance / Finance
- Congestion models (queues)
- Environmental applications
- Search problems
- Reliability models
- Statistics

Why are Rare Events Difficult to Assess?

- Typically no closed forms (complex systems)
- But crudely implemented simulation might not be good



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Why are Rare Events Difficult to Assess?

- **Relative mean squared error (RMSE) PER TRIAL** = stdev / mean

$$\frac{\sqrt{P(\text{RED})(1 - P(\text{RED}))}}{P(\text{RED})} \approx \frac{1}{\sqrt{P(\text{RED})}}$$

- *General Focus*: Estimate $P(A)$ assuming $P(A) \approx 0$.

Performance Analysis

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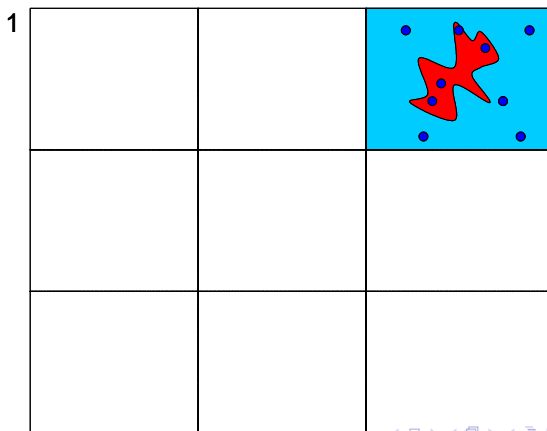
- **WEAK EFFICIENCY**: For each $\varepsilon > 0$

$$\text{RMSE for } P(A) = O(1/P(A)^\varepsilon)$$

Graphical Interpretation of Importance Sampling

- **Importance sampling (I.S.):** sample from the important region and correct via likelihood ratio

$$\text{RED AREA} \approx \text{PROPORTION DARTS IN RED AREA} \times \frac{1}{9}$$



Importance Sampling

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- **NOTE:** \tilde{P} is called a change-of-measure

- Suppose we choose $\tilde{P}(\cdot) = P(\cdot|A)$

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- **Lesson:** *Try choosing $\tilde{P}(\cdot)$ close to $P(\cdot|A)$!*

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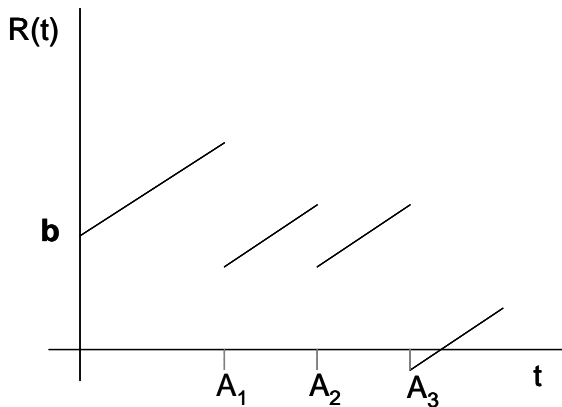
$$R(t) = b + pt - \sum_{j=1}^{N(t)} V_j$$

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$$R(t) = b + pt - \sum_{j=1}^{N(t)} V_j$$

- $N(t) = \#$ arrivals up to time t

Plot of risk reserve



- Evaluating reserve at arrival times we get random walk

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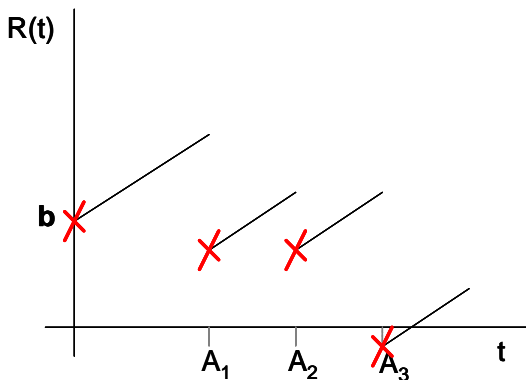
$$S(n) = b + Y_1 + \dots + Y_n$$

Insurance

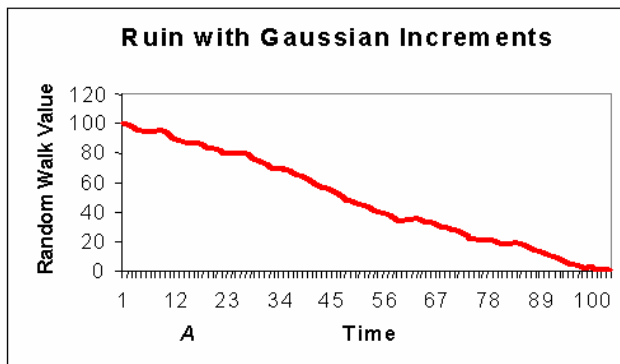
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- Suppose Y_1, Y_2, \dots are i.i.d.

$$S(n) = b + Y_1 + \dots + Y_n$$

- $R(A_n) = S(n)$ reserve at arrival times with $Y_n = p\tau_n - V_n$.

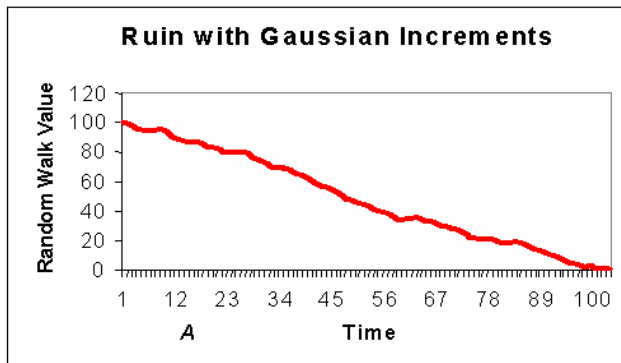


Insurance Process Conditioned on Ruin



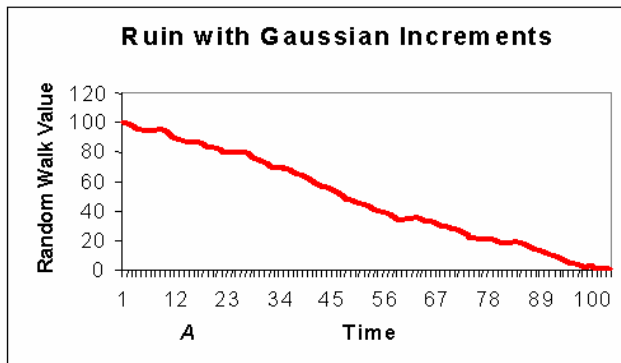
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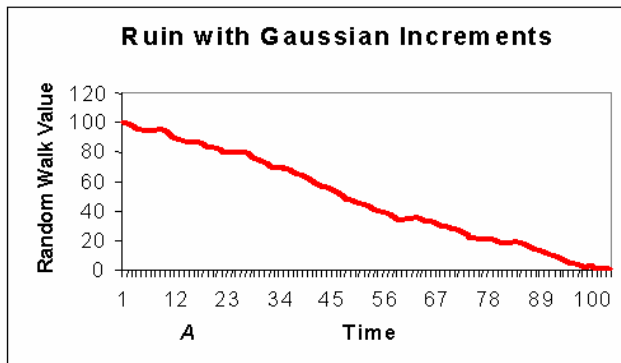
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- Light tails: Exponential, Gamma, Gaussian, mixtures of these, etc.
- Picture generated with Siegmund's 76 algorithm

- In light-tailed cases there is large deviations theory (ref. Dembo and Zeitouni '99).

Light Tails Setting: Asymptotic Conditional Distributions

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- Large deviations allows to obtain as $b \nearrow \infty$

$$P(Y_1 \leq x, \dots, Y_k \leq x | \text{ruin starting at } b) \approx \tilde{P}(Y_1 \leq x) \dots \tilde{P}(Y_k \leq x),$$

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- Suggested change-of-measure: Sample Y_k 's i.i.d. using $\tilde{P}(\cdot)$

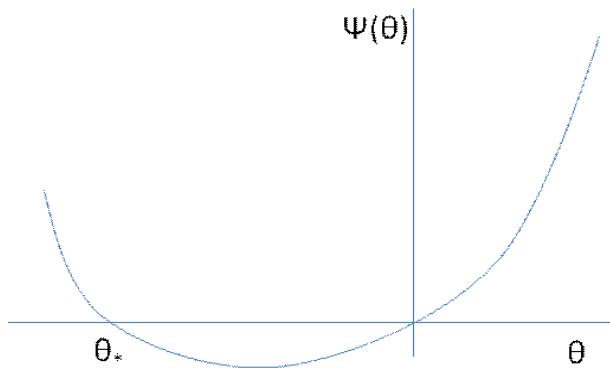
$$L = \frac{p(Y_1)}{\tilde{p}(Y_1)} \cdot \frac{p(Y_2)}{\tilde{p}(Y_2)} \cdot \dots \cdot \frac{p(Y_{\text{ruin time}})}{\tilde{p}(Y_{\text{ruin time}})}.$$

Light Tails Setting: Exponential Tilting

- More precisely if $p(\cdot)$ is the density of Y_i

$$\tilde{p}(y) = p(y) \exp(\theta_* y)$$

where $\theta_* < 0$ solves $\psi(\theta_*) = \log E \exp(\theta_* Y_i) = 0$.



Siegmund's Algorithm for Random Walks

Theorem (Siegmund '76)

Assume $\psi(\theta_* - \delta) < \infty$ for some $\delta > 0$. Then, the estimator

$$\begin{aligned} L &= \frac{p(Y_1)}{\tilde{p}(Y_1)} \cdot \frac{p(Y_2)}{\tilde{p}(Y_2)} \cdot \dots \cdot \frac{p(Y_{\text{ruin time}})}{\tilde{p}(Y_{\text{ruin time}})} \\ &= \exp(-\theta_* [Y_1 + \dots + Y_{\text{ruin time}}]) \end{aligned}$$

is **STRONGLY EFFICIENT**. Moreover, $\tilde{p}(\cdot)$ is the **ONLY STATE-INDEPENDENT** change-of-measure that achieves efficiency.

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Moral of the Story for Light Tails...

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- 3 Items 1) and 2) provide systematic tools for rare-event simulation

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Description of Systematic Approach

- State-dependent random walk: $t \in \{0, \Delta, 2\Delta, \dots\}$

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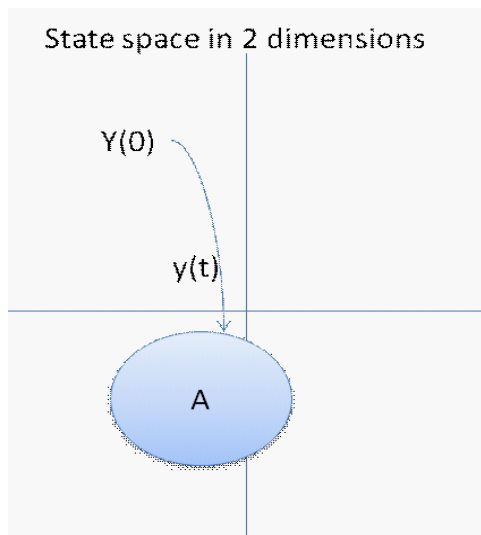
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- As $\Delta \rightarrow 0$ under mild assumptions $Y. \rightarrow y(\cdot)$ so that

$$\dot{y}(t) = E[X(y(t))].$$

Fluid Limit of State-dependent Random Walk



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- Given $z(t)$ described by an ODE so that $z(t) \neq y(t)$ one often has

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- Associated action integral

$$J(z) = \int_0^{\infty} I(\dot{z}(t), z(t)) dt.$$

Description of Systematic Approach

- A generic rare-event estimation problem:

$$\begin{aligned} & \Delta \log P(\text{Hit } B \text{ prior to } A) \\ \approx & - \inf \{ J(z) : z(\cdot) \text{ is path that hits } B \text{ prior to } A \} \end{aligned}$$

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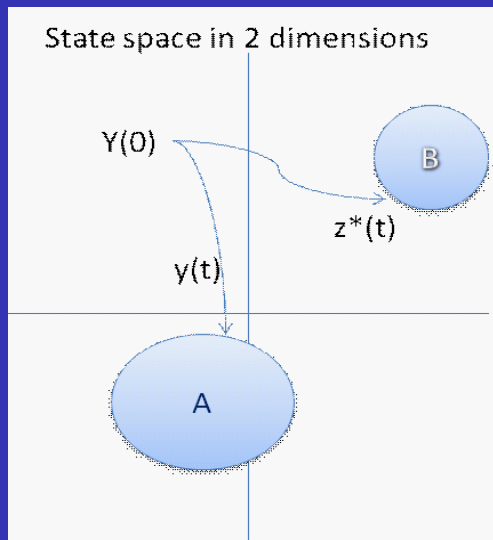
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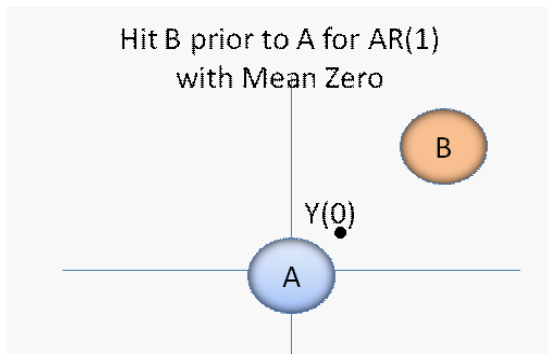
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- **Importance Sampling Strategy:**

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- 2 Follow path to approximate $z^*(t)$...

Example: Two dimensional AR(1) Process

- **Model:** $Y_{t+\Delta} - Y_t = -\Delta Y_t + \Delta X_{t+\Delta}$, X_t 's i.i.d. $N(0, I)$ and $y_0 = (1, 1)$.
- **Estimate:** $P_{y_0}(T_B < T_A)$, with
 $B = \{x : \left\|x - (e + 1/\sqrt{2}, e + 1/\sqrt{2})\right\|_2 \leq 1\}$ &
 $A = \{x : \|x\|_2 \leq 1\}$.



Example: Computing the Optimal Path

- **Calculus of Variations Problem:**

$$\min_{z \in C} \frac{1}{2} \int_0^T \|\dot{z}(t) + z(t)\|_2^2 dt$$

where

$$C = \{z : z(0) = (1, 1), \left\| z(T) - (e + 2^{-1/2})(1, 1) \right\|_2 \leq 1, \\ T < \infty, \|z(t)\|_2 > 1, t < T\}$$

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- **Solution:**

$$\begin{aligned} z(t) &= (\exp(t), \exp(t)) \\ \dot{z}(t) &= z(t). \end{aligned}$$

- **Importance Sampling** Q : W_t is i.i.d. $N(0,1)$ under Q

$$Y_{t+\Delta} - Y_t = \dot{z}(t) \Delta + \Delta W_{t+\Delta} \implies \text{Fluid } dy_t \approx \dot{z}(t) dt$$

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Interpreting the Change-of-Measure

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- **Likelihood ratio** dP/dQ :

$$L_0 = \exp \left(- \sum_{j=0}^{T_B/\Delta-1} (\dot{z}(j\Delta) + Y_{j\Delta}) \cdot X_{(j+1)\Delta} + \sum_{j=0}^{T_B/\Delta-1} \frac{\|\dot{z}(j\Delta) + Y_{j\Delta}\|_2^2}{2} \right)$$

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$$\text{IS Estimator} = L_0 \times I(T_B < T_A)$$

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- **Importance Sampling** Q : W_t is Brownian motion under Q

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- **First note:** $Y_{t+\Delta} - Y_t = -\Delta Y_t + \Delta X_{t+\Delta} \implies$

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- **Get likelihood ratio representation:**

$$\begin{aligned} L &= \exp \left(- \sum_{j=0}^{T_B/\Delta-1} 2 Y_{(j\Delta)} \cdot X_{(j+1)\Delta} + \sum_{j=0}^{T_B/\Delta-1} 2 \|Y_{(j\Delta)}\|_2^2 \right) \\ &= \exp \left(- \|Y_{T_B}\|_2^2 / \Delta + \|y_0\|_2^2 / \Delta + \Delta \sum_{j=0}^{T_B/\Delta-1} \|W_{(j+1)\Delta} + Y_{j\Delta}\|_2^2 \right) \end{aligned}$$

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- **Second moment of estimator:**

$$E^Q[L^2 \times I(T_B < T_A)] = \exp \left(-2 \left(\inf_{y \in B} \|y\|_2^2 - \|y_0\|_2^2 \right) / \Delta + o(1/\Delta) \right)$$

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- **Asymptotic Optimality follows since:**

$$P_{y_0}(T_B < T_A) = \exp \left(- \left(\inf_{y \in B} \|y\|_2^2 - \|y_0\|_2^2 \right) / \Delta + o(1/\Delta) \right).$$

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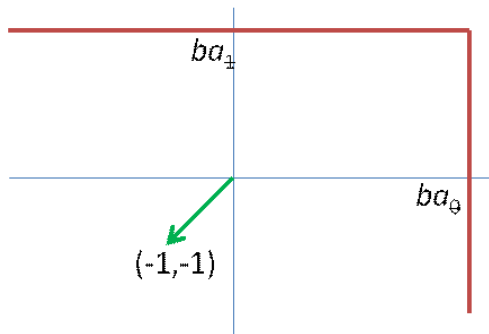
Counter-examples

- **Model:** $Y_t = (-1, -1)t + X_t$; X_t is Brownian motion &
 $B_b = \{(x, y) : x \geq a_0 b \text{ or } y \geq a_1 b\}$

Counter-examples

- **Model:** $Y_t = (-1, -1)t + X_t$; X_t is Brownian motion & $B_b = \{(x, y) : x \geq a_0 b \text{ or } y \geq a_1 b\}$
- **Estimate:** $u(b) = P(T_{B_b} < \infty)$ as $b \nearrow \infty$

First Passage Time Problem in two dimensions



- **Known result:** As $b \nearrow \infty$

$$u(b) = \exp(-2b \min\{a_0, a_1\}) (1 + o(1)).$$

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- **Let us assume** $a_0 < a_1 \dots$

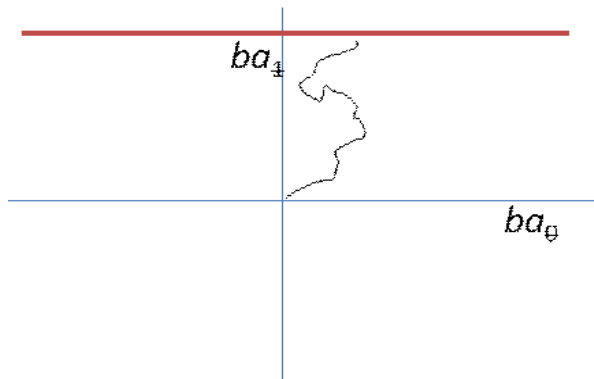
- **Change-of-measure Q_0 :** Brownian Motion with drift $(1, -1)$.

Second Moment of Estimator:

$$\begin{aligned} \text{2nd Moment} &= E^Q[\exp\left(- (4, 0) \cdot X_{T_{B_b}}\right) I(T_{B_b} < \infty)] \\ &= E[\exp\left(- (2, 0) X_{T_{B_b}}\right) I(T_{B_b} < \infty)]. \end{aligned}$$

- What happens if a path exits $A_b = \{(x, y) : y \geq a_1 b\}$?

Hitting the Unlikely Side...



- **Lower bound:**

$$\begin{aligned} \text{2nd Mnt} &\geq E[\exp(-2X_{T_{A_b}}^{(1)}) I(T_{B_b} < \infty, T_{A_b} = T_{B_b})] \\ &= E^{Q_1}[\exp(-2X_{T_{A_b}}^{(1)} - 2a_1 b) I(T_{A_b} = T_{B_b})], \end{aligned}$$

where Q_1 is change-of-measure yielding Brownian Motion with drift $(-1, 1)$.

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- Note that $X_{T_{A_b}}^{(1)} \approx -a_0 a_1 b$ and $T_{A_b} = T_{B_b}$ under Q_1
- **Consequently:** 2nd Moment at least roughly

$$\exp(-2a_1 b(1 - a_0))$$

Can easily pick $a_0 < a_1$ to **break asymptotic optimality and even get HUGE variance!!**

- **Model: Random walk...**

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for $\alpha > 0$ and $L(t\beta) / L(t) \rightarrow 1$ as $t \rightarrow \infty$ for $\beta > 0$ (e.g. $L(t) = \log(1+t)$).

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- Note $E \exp(\theta X_1) = \infty$ for $\theta > 0$.

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- **No clear way to mimic zero variance change-of-measure!**
- **No clear way to apply the systematic approach!**

Solution? State-dependent Importance Sampling

As we shall see these examples can be addressed using

State – dependent importance sampling

which we will study in the second part of this lecture...

Summary and Conclusions

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- Optimal path in large deviations dictates tilting (several ways of interpreting)
- Approach fails (badly!) in non-convex problems and in heavy-tailed situations