## Testing Distribution Properties

## Constantinos (Costis) Daskalakis

CSAIL and EECS, MIT


Jayadev Acharya Cornell


Nishanth Dikkala CSAIL, MIT


Gautam Kamath CSAIL, MIT

## BIG Data



Facebook: $\mathbf{2 0}$ petabytes images daily


Human genome: 40 exabytes storage by 2025


SKA Telescope: 1 exabyte daily

## B/G Data



Facebook: $\mathbf{2 0}$ petabytes images daily


Human genome: 40 exabytes storage by 2025


SKA Telescope: 1 exabyte daily

## High-dimensional

## Expensive




Computer vision

Experimental drugs

## What properties do your BIG distributions have?



## e.g.1: play the lottery?



## e.g.1: play the lottery?



## e.g. 1.1: Polish MultiLotek



## e.g. 1.1: Polish MultiLotek

- Chooses "uniformly" at random distinct 20 numbers in $\{1, \ldots, 80\}$.



## e.g. 1.1: Polish MultiLotek

- Chooses "uniformly" at random distinct 20 numbers in $\{1, \ldots, 80\}$.
- Initial machine biased



Thanks to Krzysztof Onak (pointer) and Eric Price (graph)

## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k (=3,4) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.


## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k ( =3,4) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.
- Data:
- Pick 3-8522 results from 5/22/75 to 10/15/00


## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k ( =3,4) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.
- Data:
- Pick 3-8522 results from 5/22/75 to 10/15/00
- $\chi^{2}$-test (on Excel) answers " $42 \%$ confidence"


## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k ( $=3,4$ ) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.
- Data:
- Pick 3-8522 results from 5/22/75 to 10/15/00
- $\chi^{2}$-test (on Excel) answers " $42 \%$ confidence"
- Pick 4-6544 results from 9/1/77 to 10/15/00.


## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k (=3,4) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.
- Data:
- Pick 3-8522 results from 5/22/75 to 10/15/00
- $\chi^{2}$-test (on Excel) answers " $42 \%$ confidence"
- Pick 4-6544 results from 9/1/77 to 10/15/00.
- fewer results than possible values
- not a good idea to run $\chi^{2}$ test


## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k (=3,4) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.
- Data:
- Pick 3-8522 results from 5/22/75 to 10/15/00
- $\chi^{2}$-test (on Excel) answers " $42 \%$ confidence"
- Pick 4-6544 results from 9/1/77 to 10/15/00.
- fewer results than possible values
- not a good idea to run $\chi^{2}$ test
- convergence to $\chi^{2}$ distribution won't kick in for small sample size


## e.g. 1.2: New Jersey Pick 3,4 Lottery

- New Jersey Pick k (=3,4) Lottery.
- Pick $k$ random digits in order.
$-10^{k}$ possible values.
- Data:
- Pick 3-8522 results from 5/22/75 to 10/15/00
- $\chi^{2}$-test (on Excel) answers " $42 \%$ confidence"
- Pick 4-6544 results from 9/1/77 to 10/15/00.
- fewer results than possible values
- not a good idea to run $\chi^{2}$ test
- convergence to $\chi^{2}$ distribution won't kick in for small sample size
- (textbook) rule of thumb: expected number of at least 5 observations of each element in the domain under the null hypothesis to run $\chi^{2}$


## e.g.2: Linkage Disequilibrium



Single nucleotide polymorphisms, are they independent?

## e.g.2: Linkage Disequilibrium

Genome


Single nucleotide polymorphisms, are they independent?
Suppose $n$ loci, 2 possible states each, then:

## e.g.2: Linkage Disequilibrium

Genome


Single nucleotide polymorphisms, are they independent?
Suppose $n$ loci, 2 possible states each, then:

- state of one's genome $\in\{0,1\}^{n}$


## e.g.2: Linkage Disequilibrium

Genome


Single nucleotide polymorphisms, are they independent?
Suppose $n$ loci, 2 possible states each, then:

- state of one's genome $\in\{0,1\}^{n}$
- humans: some distribution $p$ over $\{0,1\}^{n}$


## e.g.2: Linkage Disequilibrium

Genome


Single nucleotide polymorphisms, are they independent?
Suppose $n$ loci, 2 possible states each, then:

- state of one's genome $\in\{0,1\}^{n}$
- humans: some distribution $p$ over $\{0,1\}^{n}$

Question: Is $p$ a product dist'n OR far from all product dist'ns?

## e.g.2: Linkage Disequilibrium

Genome


Single nucleotide polymorphisms, are they independent?
Suppose $n$ loci, 2 possible states each, then:

- state of one's genome $\in\{0,1\}^{n}$
- humans: some distribution $p$ over $\{0,1\}^{n}$

Question: Is $p$ a product dist'n OR far from all product dist'ns?
should we expect the genomes from the 1000 human genomes project to be sufficient? up to how many loci?

## e.g. 3: Outbreak of diseases

- Similar patterns in different years?
- More prevalent near large airports?

- Flu 2005
- Flu 2006


# Old questions, new challenges 

Classical Setting

Modern Setting

## Old questions, new challenges

Classical Setting

Domain:


Domain:


One human genome

## Old questions, new challenges

## Classical Setting

Domain:


1000 tosses

Small domain $D$
n large, $|D|$ small
(comparatively)

Domain:

One human genome

Large domain $D$
n small, $|D|$ large

## Old questions, new challenges

## Classical Setting

Domain:


1000 tosses

Small domain $D$
n large, $|D|$ small
(comparatively)

Asymptotic analysis
Computation not crucial

Domain:


One human genome

Large domain $D$
n small, $|D|$ large

## Old questions, new challenges

## Classical Setting

Domain:


1000 tosses

Small domain $D$
n large, $|D|$ small
(comparatively)
Asymptotic analysis
Computation not crucial

Domain:


One human genome

Large domain $D$
n small, $|D|$ large

New challenges:
samples, computation, communication, storage

## A Key Question

- How many samples do you need in terms of domain size?


## A Key Question

- How many samples do you need in terms of domain size?
- Do you need to estimate the probabilities of each domain item?
-- OR --
- Can sample complexity be sublinear in size of the domain?


## A Key Question

- How many samples do you need in terms of domain size?
- Do you need to estimate the probabilities of each domain item?
-- OR --
- Can sample complexity be sublinear in size of the domain?

Rules out standard statistical techniques

## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## Problem formulation

## Model

$\mathcal{P}$ : family of distributions over $D$

## Problem formulation

## Model

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem formulation

## Model

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

$$
p \in \mathcal{P} \quad \text { vs } \quad d(p, \mathcal{P})>\varepsilon
$$

## Problem formulation

## Model

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

$$
p \in \mathcal{P} \quad \text { vs } \quad d(p, \mathcal{P})>\varepsilon
$$

## Objective

Minimize samples
Computational efficiency

## Problem formulation

## Model

## discrete

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

$$
p \in \mathcal{P} \quad \text { vs } \quad d(p, \mathcal{P})>\varepsilon
$$

## Objective

Minimize samples
Computational efficiency

## Problem formulation

## Model

## discrete

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

## Objective

Minimize samples

$$
\begin{aligned}
& p \in \mathcal{P} \quad \text { vs } \overbrace{\min _{q \in \mathcal{P}} \frac{\ell_{1}(p, q)}{2}}^{d(p, \mathcal{P})>\varepsilon} \\
& \text { les }
\end{aligned}
$$

Computational efficiency

## Problem formulation

## Model

## discrete

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

## Objective

Minimize samples
Computational efficiency

$$
\max _{\text {events } \mathcal{E}}|p(\mathcal{E})-q(\mathcal{E})| \equiv d_{T V}(p, q)
$$

## Problem formulation

## Model

```
discrete
```

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

## Objective

Minimize samples
Computational efficiency


## Problem formulation

## Model

$\mathcal{P}$ : family of distributions over $D$
may be non-parametric, e.g. unimodal, product, log-concave

## Problem

Given: samples from unknown $p$
with probability 0.9 , distinguish

## Objective

Minimize samples
Computational efficiency


## Well-studied Problem

(Composite) hypothesis testing

- Neyman-Pearson test
- Kolmogorov-Smirnov test
- Pearson's chi-squared test
- Generalized likelihood ratio test
- ...


## Well-studied Problem

(Composite) hypothesis testing

- Neyman-Pearson test
- Kolmogorov-Smirnov test
- Pearson’s chi-squared test
- Generalized likelihood ratio test


## Quantities of Interest

$$
\begin{aligned}
P_{F} & =\operatorname{Pr}(\text { accept when hypothesis false }) \\
P_{M} & =\operatorname{Pr}(\text { reject when hypothesis true })
\end{aligned}
$$

## Well-studied Problem

(Composite) hypothesis testing

- Neyman-Pearson test
- Kolmogorov-Smirnov test
- Pearson's chi-squared test
- Generalized likelihood ratio test


## Quantities of Interest

$$
\begin{aligned}
& P_{F}=\operatorname{Pr}(\text { accept when hypothesis false }) \\
& P_{M}=\operatorname{Pr}(\text { reject when hypothesis true })
\end{aligned}
$$

## Focus

Consistency
Error exponents: $\exp (-s \cdot R)$ as $s \rightarrow \infty$

## Well-studied Problem

(Composite) hypothesis testing

- Neyman-Pearson test
- Kolmogorov-Smirnov test
- Pearson's chi-squared test
- Generalized likelihood ratio test


## Quantities of Interest

$$
\begin{aligned}
P_{F} & =\operatorname{Pr}(\text { accept when hypothesis false }) \\
P_{M} & =\operatorname{Pr}(\text { reject when hypothesis true })
\end{aligned}
$$

## Focus

Consistency
Error exponents: $\exp (-s \cdot R)$ as $s \rightarrow \infty$
Asymptotic regime: Results kick in when $s \gg|D|$

Why Most Published Research Findings Are False
John P. A. Ioannidis
Published: August 30, 2005 • http://dx.doi.org/10.1371/journal.pmed. 0020124

## oblem

| Article | Authors | Metrics | Comments | Related Content |
| :--- | :--- | :--- | :--- | :--- |
| $\approx \approx$ |  |  |  |  |

Abstract
Modeling the Framework
for False Positive
Findings
Bias
Testing by Several Independent Teams

Corollaries
Most Research Findings
Are False for Most Research Designs and for Most Fields

Claimed Research Findings May Often Be Simply Accurate Measures of the Prevailing Bias

Abstract

## Summary

There is increasing concern that most current published research findings are false. The probability that a research claim is true may depend on study power and bias, the number of other studies on the same question, and, importantly, the ratio of true to no relationships among the relationships probed in each scientific field. In this framework, a research finding is less likely to be true when the studies conducted in a field are smaller; when effect sizes are smaller; when there is a greater number and lesser preselection of tested relationships; where there is greater flexibility in designs, definitions, outcomes, and analytical modes; when there is greater financial and other interest and prejudice; and when more teams are involved in a scientific field in chase of statistical significance. Simulations show that for most study designs and settings, it is more likely for a research claim to be false than true. Moreover, for many current scientific fields, claimed research findings may often be simply accurate measures of the prevailing bias. In this essay, I discuss the implications of these problems for the conduct and interpretation of research.
thesis false)
$P_{M}=\operatorname{Pr}($ reject when hypothesis true)

## Focus

Consistency
Error exponents: $\exp (-s \cdot R)$ as $s \rightarrow \infty$ Asymptotic regime: Results kick in when $s \gg|D|$

# oblem 

Article
$\approx$

## Study delivers bleak verdict on validity of psychology experiment results

Abstract
Modeling the Framews for False Positive
Findings
Bias
Testing by Several Independent Teams Corollaries

Most Research Findin
Are False for Most
Research Designs ans for Most Fields

Claimed Research Findings May Often B Simply Accurate Measures of the Prevailing Bias

## Focu

Of 100 studies published in top-ranking journals in 2008, $75 \%$ of social psychology experiments and half of cognitive studies failed the replication test Psychology experiments are failing the replication test - for good reason

esis false)
esis true)

C

- There are many reasons why an experiment might fail to replicate, but more than this, the study has highlighted some issues with academic publishing and modern science. Photograph: Pere Sanz / Alamy/Alamy

A major investigation into scores of claims made in psychology research journals has delivered a bleak verdict on the state of the science.

Asymptotic regime: Results kick in when $s \gg|D|$

## oblem

Article
$\approx$

## Study delivers bleak verdict on validity of psychology experiment results

Abstract
Modeling the Framews for False Positive Findings
Bias
Testing by Several Independent Teams Corollaries

Most Research Findin
Are False for Most
Research Designs anc for Most Fields

Claimed Research Findings May Often B Simply Accurate Measures of the Prevailing Bias

## Focu

Of 100 studies published in top-ranking journals in 2008, $75 \%$ of social psychology experiments and half of cognitive studies failed the replication test Psychology experiments are failing the replication test - for good reason

esis false) esis true)

- Sublinear
in $|D|$ ?
- There are many reasons why an experiment might fail to replicate, but more than this, the study has highlighted some issues with academic publishing and modern science. Photograph: Pere Sanz / Alamy/Alamy

A major investigation into scores of claims made in psychology research journals has delivered a bleak verdict on the state of the science.

Asymptotic regime: Results kick in when $s \gg|D|$

Why Most Published Research Findings Are False John P. A. loannidis

Published: August 30, 2005 • http://dx.doi.org/10.1371/journal.pmed. 0020124

## oblem

$\square$ Study delivers bleak verdict on validity of psychology experiment results

Abstract
Modeling the Framews for False Positive Findings
Bias
Testing by Several Independent Teams Corollaries

Most Research Findin
Are False for Most
Research Designs anc for Most Fields

Claimed Research Findings May Often B Simply Accurate Measures of the Prevailing Bias

Focu:
Of 100 studies published in top-ranking journals in 2008, $75 \%$ of social psychology experiments and half of cognitive studies failed the replication test Psychology experiments are failing the replication test - for good reason


- There are many reasons why an experiment might fail to replicate, but more than this, the study has highlighted some issues with academic publishing and modern science. Photograph: Pere Sanz / Alamy/Alamy

A major investigation into scores of claims made in psychology research journals
$E$
has delivered a bleak verdict on the state of the science.
1 -
Asymptotic regime: Results kick in when $s \gg|D|$
esis false)
esis true)

- Sublinear
in $|D|$ ?
- Strong control for false negatives?


## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## The Menu



## Testing the Fairness of a Coin

- b : unknown probability of a
- Question: Is $b=0.5$ OR $|b-0.5| \geq \epsilon$ ?
- Goal: Toss coin several times, deduce correct answer w/prob. $\geq 0.99$


## Testing the Fairness of a Coin

- b : unknown probability of a

- Question: Is $b=0.5 \quad$ OR $\quad|b-0.5| \geq \epsilon$ ?
- Goal: Toss coin several times, deduce correct answer w/ prob. $\geq 0.99$


## Testing the Fairness of a Coin

- b : unknown probability of a

- Question: Is $b=0.5 \quad$ OR $\quad|b-0.5| \geq \epsilon$ ?
- Goal: Toss coin several times, deduce correct answer w/ prob. $\geq 0.99$
- Number of samples required?
- Tight answer $\Theta\left(\frac{1}{\epsilon^{2}}\right)$


## Testing the Fairness of a Coin

- b : unknown probability of a

- Question: Is $b=0.5 \quad$ OR $\quad|b-0.5| \geq \epsilon$ ?
- Goal: Toss coin several times, deduce correct answer w/ prob. $\geq 0.99$
- Number of samples required?
- Tight answer $\Theta\left(\frac{1}{\epsilon^{2}}\right)$
- Upper bound: compare average to 0.5 , reject if farther than $\frac{\epsilon}{2}$


## Testing the Fairness of a Coin

- b : unknown probability of a

- Question: Is $\quad b=0.5 \quad$ OR $\quad|b-0.5| \geq \epsilon$ ?
- Goal: Toss coin several times, deduce correct answer w/ prob. $\geq 0.99$
- Number of samples required?
- Tight answer $\Theta\left(\frac{1}{\epsilon^{2}}\right)$
- Upper bound: compare average to 0.5 , reject if farther than $\frac{\epsilon}{2}$
- Lower bound: a sleek one uses the subadditivity of Hellinger ${ }^{2}$ distance


## Testing Uniformity

- $p$ : unknown distribution over $D$
- sample access to $p$
- Question: is $p=U_{D}$ or $d_{\mathrm{TV}}\left(p, U_{D}\right)>\epsilon$ ?


## Testing Uniformity

- $p$ : unknown distribution over $D$
- sample access to $p$
- $\mathcal{P}=\left\{\right.$ Uniform $\left._{D}\right\}$
- $p$ : unknown
$p \in \mathcal{P} \quad$ vs $\quad d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=U_{D}$ or $d_{\mathrm{TV}}\left(p, U_{D}\right)>\epsilon$ ?


## Testing Uniformity

- $p$ : unknown distribution over $D$
- sample access to $p$
- $\mathcal{P}=\left\{\right.$ Uniform $\left._{D}\right\}$
- $p$ : unknown
$p \in \mathcal{P} \quad$ vs $\quad d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=U_{D}$ or $d_{\mathrm{TV}}\left(p, U_{D}\right)>\epsilon$ ?
- [Paninski'03]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time


## Testing Uniformity

- $p$ : unknown distribution over $D$
- sample access to $p$
- $\mathcal{P}=\left\{\right.$ Uniform $\left._{D}\right\}$
- $p$ : unknown
$p \in \mathcal{P} \quad$ vs $\quad d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=U_{D}$ or $d_{\mathrm{TV}}\left(p, U_{D}\right)>\epsilon$ ?
- [Paninski'03]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time
"Intuition:"
- Lower Bound: Suppose $q$ is uniform distribution over $\{1, \ldots, m\}$ and $p$ is uniform on random $m / 2$ size subset of $\{1, \ldots, m\}$


## Testing Uniformity

- $p$ : unknown distribution over $D$
- sample access to $p$
- $\mathcal{P}=\left\{\right.$ Uniform $\left._{D}\right\}$
$p$ : unknown
$p \in \mathcal{P} \quad$ vs $\quad d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=U_{D}$ or $d_{\mathrm{TV}}\left(p, U_{D}\right)>\epsilon$ ?
- [Paninski'03]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time


## "Intuition:"

- Lower Bound: Suppose $q$ is uniform distribution over $\{1, \ldots, m\}$ and $p$ is uniform on random $m / 2$ size subset of $\{1, \ldots, m\}$
- Unless $\Omega(\sqrt{m})$ samples are observed, there are no collisions, hence cannot distinguish between $q$ or $p$ chosen as above


## Testing Uniformity

- $p$ : unknown distribution over $D$
- sample access to $p$
- $\mathcal{P}=\left\{\right.$ Uniform $\left._{D}\right\}$
$p$ : unknown $p \in \mathcal{P} \quad$ vs
$d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=U_{D}$ or $d_{\mathrm{TV}}\left(p, U_{D}\right)>\epsilon$ ?
- [Paninski'03]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time


## "Intuition:"

- Lower Bound: Suppose $q$ is uniform distribution over $\{1, \ldots, m\}$ and $p$ is uniform on random $m / 2$ size subset of $\{1, \ldots, m\}$
- Unless $\Omega(\sqrt{m})$ samples are observed, there are no collisions, hence cannot distinguish between $q$ or $p$ chosen as above
- Upper Bound: Collision statistics suffice to distinguish


## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## Identity Testing ("goodness of fit")

- $p, q$ : distributions over $D$
- $q$ : given; sample access to $p$
- Question: is $p=q$ or $d_{\mathrm{TV}}(p, q)>\epsilon$ ?


## Identity Testing ("goodness of fit")

- $p, q$ : distributions over $D$
- $\quad$ : given; sample access to $p$
- $\mathcal{P}=\{q\}$
- $p:$ unknown
$p \in \mathcal{P} \quad$ vs $\quad d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=q$ or $d_{\mathrm{TV}}(p, q)>\epsilon$ ?


## Identity Testing ("goodness of fit")

- $p, q$ : distributions over $D$
- $q$ : given; sample access to $p$
- $\mathcal{P}=\{q\}$
$p$ : unknown
$p \in \mathcal{P} \quad$ vs
$d_{\mathrm{TV}}(p, \mathcal{P})>\varepsilon$
- Question: is $p=q$ or $d_{\mathrm{TV}}(p, q)>\epsilon$ ?
- [Batu-Fisher-Fortnow-Kumar-Rubinfeld-White’01]...
- [Paninski'08, Valiant-Valiant'14]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time


## Identity Testing ("goodness of fit")

- $p, q$ : distributions over $D$
- $q$ : given; sample access to $p$

- Question: is $p=q$ or $d_{\mathrm{TV}}(p, q)>\epsilon$ ?
- [Batu-Fisher-Fortnow-Kumar-Rubinfeld-White’01]...
- [Paninski'08, Valiant-Valiant'14]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time
- [w/ Acharya-Kamath NIPS'15]: a tolerant goodness of fit test with same sample size can distinguish: $\chi^{2}(p, q) \leq \epsilon^{2} / 2 \quad$ vs $\quad \ell_{1}^{2}(p, q)>\epsilon^{2}$


## Identity Testing ("goodness of fit")

- $p, q$ : distributions over $D$
- $\quad$ : given; sample access to $p$

- Question: is $p=q$ or $d_{\mathrm{TV}}(p, q)>\epsilon$ ?
- [Batu-Fisher-Fortnow-Kumar-Rubinfeld-White’01]...
- [Paninski'08, Valiant-Valiant'14]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time
- [w/ Acharya-Kamath NIPS'15]: a tolerant goodness of fit test with same sample size can distinguish: $\chi^{2}(p, q) \leq \epsilon^{2} / 2$ vs $\ell_{1}^{2}(p, q)>\epsilon^{2}$
- $\chi^{2}(p, q):=\sum_{i \in D} \frac{\left(p_{i}-q_{i}\right)^{2}}{q_{i}}$
- Cauchy-Schwarz: $\chi^{2}(p, q) \geq \ell_{1}(p, q)^{2}$


## Identity Testing ("gor

- $p, q$ : distributions over $D$
- $q$ : given; sample access to $p$
- Question: is $p=q$ or $d_{\mathrm{TV}}(p, q)>\epsilon$ ?
- [Batu-Fisher-Fortnow-Kumar-Rubinfeld-U
- [Paninski'08, Valiant-Valiant'14]: $\Theta\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time
- [w/ Acharya-Kamath NIPS'15]: a tolerant goodness of fit test with same sample size can distinguish: $\chi^{2}(p, q) \leq \epsilon^{2} / 2$ vs $\ell_{1}^{2}(p, q)>\epsilon^{2}$
- $\chi^{2}(p, q):=\sum_{i \in D} \frac{\left(p_{i}-q_{i}\right)^{2}}{q_{i}}$
- Cauchy-Schwarz: $\chi^{2}(p, q) \geq \ell_{1}(p, q)^{2}$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw $m$ samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw $m$ samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw $m$ samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$

Want:
$-Z$ small $\rightarrow$ Case 1
$-Z$ large $\rightarrow$ Case 2

## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- $N_{i} \sim \operatorname{Poisson}\left(m \cdot p_{i}\right)$
- $\left(N_{i}\right)_{i \in D}$ independent random variables
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- $N_{i} \sim \operatorname{Poisson}\left(m \cdot p_{i}\right)$
- $\left(N_{i}\right)_{i \in D}$ independent random variables
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$
$-E[Z]=m \cdot \chi^{2}(p, q)$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- $N_{i} \sim \operatorname{Poisson}\left(m \cdot p_{i}\right)$
- $\left(N_{i}\right)_{i \in D}$ independent random variables
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$
$-E[Z]=m \cdot \chi^{2}(p, q)$
- Case 1: $E[Z] \leq m \cdot \epsilon^{2} / 2$; Case 2: $E[Z] \geq m \cdot \epsilon^{2}$


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- $N_{i} \sim \operatorname{Poisson}\left(m \cdot p_{i}\right)$
- $\left(N_{i}\right)_{i \in D}$ independent random variables
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$
$-E[Z]=m \cdot \chi^{2}(p, q)$
- Case 1: $E[Z] \leq m \cdot \epsilon^{2} / 2$; Case 2: $E[Z] \geq m \cdot \epsilon^{2}$
- chug chug chug...bound variance of $Z$
$\rightarrow \mathrm{O}\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples suffice to distinguish


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- $N_{i} \sim \operatorname{Poisson}\left(m \cdot p_{i}\right)$
- $\left(N_{i}\right)_{i \in D}$ independent random variables


## Side-Note:

- Pearson's $\chi^{2}$ test uses statistic $\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}}{m \cdot q_{i}}$
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$
$-E[Z]=m \cdot \chi^{2}(p, q)$
- Case 1: $E[Z] \leq m \cdot \epsilon^{2} / 2$; Case 2: $E[Z] \geq m \cdot \epsilon^{2}$
- chug chug chug...bound variance of $Z$
$\rightarrow \mathrm{O}\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples suffice to distinguish


## An Improved $\chi^{2}$ - Test

- Goal: given $q$ and sample access to $p$ distinguish:

$$
\text { Case 1: } \chi^{2}(p, q) \leq \epsilon^{2} / 2 \text { vs Case 2: } \ell_{1}^{2}(p, q) \geq \epsilon^{2}
$$

- Approach: Draw Poisson $(m)$ many samples from $p$
- $N_{i}$ : \# of appearances of symbol $i \in D$
- $N_{i} \sim \operatorname{Poisson}\left(m \cdot p_{i}\right)$
- $\left(N_{i}\right)_{i \in D}$ independent random variables
- Statistic: $Z=\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}-N_{i}}{m \cdot q_{i}}$


## Side-Note:

- Pearson's $\chi^{2}$ test uses statistic $\sum_{i} \frac{\left(N_{i}-m \cdot q_{i}\right)^{2}}{m \cdot q_{i}}$
- Subtracting $N_{i}$ in the numerator gives an unbiased estimator and
$-E[Z]=m \cdot \chi^{2}(p, q)$ importantly may hugely decrease variance
- Case 1: $E[Z] \leq m \cdot \epsilon^{2} / 2$; Case $2: E[Z] \geq m \cdot \epsilon^{2}$
- chug chug chug...bound variance of $Z$
$\rightarrow O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples suffice to distinguish


## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties
- monotonicity: "PDF is monotone," e.g. cancer vs radiation


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties
- monotonicity: "PDF is monotone," e.g. cancer vs radiation
- unimodality: "PDF is single-peaked," e.g. single source of disease


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties
- monotonicity: "PDF is monotone," e.g. cancer vs radiation
- unimodality: "PDF is single-peaked," e.g. single source of disease
- log-concavity: "log PDF is concave"


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties
- monotonicity: "PDF is monotone," e.g. cancer vs radiation
- unimodality: "PDF is single-peaked," e.g. single source of disease
- log-concavity: "log PDF is concave"
- monotone-hazard rate: "log ( $1-\mathrm{CDF}$ ) is concave"


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties
- monotonicity: "PDF is monotone," e.g. cancer vs radiation
- unimodality: "PDF is single-peaked," e.g. single source of disease
- log-concavity: "log PDF is concave"
- monotone-hazard rate: "log (1-CDF) is concave"
- product distribution, e.g. testing linkage disequilibrium


## Testing Properties of Distributions

- so far $\mathcal{P}=\{$ single distribution $\}$
- restrictive, as rarely know hypothesis distribution exactly
- natural extension: test structural properties
- monotonicity: "PDF is monotone," e.g. cancer vs radiation
- unimodality: "PDF is single-peaked," e.g. single source of disease
- log-concavity: "log PDF is concave"
- monotone-hazard rate: "log ( $1-\mathrm{CDF}$ ) is concave"
- product distribution, e.g. testing linkage disequilibrium
- Example question:
$-\mathcal{P}=\{$ unimodal distributions over $[m]\}$
- Sample access to $p$
- Is $p$ unimodal OR is $p \epsilon$-far from or unimodal distributions?


## Testing Properties of Distributions

[w/ Acharya and Kamath NIPS'15]:

1. Testing identity, monotonicity, log-concavity, monotone hazard-rate, unimodality for distributions over (ordered set) $D$ is doable w/ $O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.

## Testing Properties of Distributions

[w/ Acharya and Kamath NIPS'15]:

1. Testing identity, monotonicity, log-concavity, monotone hazard-rate, unimodality for distributions over (ordered set) $D$ is doable w/ $O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.
2. Testing monotonicity/independence of a distribution over $D=[m]^{d}$ is doable $\mathrm{w} / O\left(\frac{m^{d / 2}}{\epsilon^{2}}\right) \equiv O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.

- previous best for monotonicity testing: $\tilde{O}\left(\frac{m^{d-0.5}}{\epsilon^{4}}\right)$ [Bhattacharya-Fisher-Rubinfeld-Valiant'11]
- previous best for independence: $d=2$, worse bounds [Batu et al.'01]


## Testing Properties of Distributions

[w/ Acharya and Kamath NIPS'15]:

1. Testing identity, monotonicity, log-concavity, monotone hazard-rate, unimodality for distributions over (ordered set) $D$ is doable w/ $O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.
2. Testing monotonicity/independence of a distribution over $D=[m]^{d}$ is doable $\mathrm{w} / O\left(\frac{m^{d / 2}}{\epsilon^{2}}\right) \equiv O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.

- previous best for monotonicity testing: $\tilde{O}\left(\frac{m^{d-0.5}}{\epsilon^{4}}\right)$ [Bhattacharya-Fisher-Rubinfeld-Valiant'11]
- previous best for independence: $\mathrm{d}=2$, worse bounds [Batu et al.'01]

3. All bounds above are optimal

- i.e. matching lower bounds for both 1 and 2 via Le Cam Inequality.


## Testing Properties of Distributions

[w/ Acharya and Kamath NIPS'15]:

1. Testing identity, monotonicity, log-concavity, monotone hazard-rate, unimodality for distributions over (ordered set) $D$ is doable w/ $O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.
2. Testing monotonicity/independence of a distribution over $D=[m]^{d}$ is doable $\mathrm{w} / O\left(\frac{m^{d / 2}}{\epsilon^{2}}\right) \equiv O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.

- previous best for monotonicity testing: $\tilde{O}\left(\frac{m^{d-0.5}}{\epsilon^{4}}\right)$ [Bhattacharya-Fisher-Rubinfeld-Valiant'11]
- previous best for independence: $\mathrm{d}=2$, worse bounds [Batu et al.'01]

3. All bounds above are optimal

- i.e. matching lower bounds for both 1 and 2 via Le Cam Inequality.

4. Unified approach, computationally efficient tests, based on new $\chi^{2}$-tolerant tester

## Testing Properties of Distributions

[w/ Acharya and Kamath NIPS'15]:

1. Testing identity, monotonicity, log-concavity, monotone hazard-rate, unimodality for distributions over (ordered set) $D$ is doable w/ $O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.
2. Testing monotonicity/independence of a distribution over $D=[m]^{d}$ is doable $\mathrm{w} / O\left(\frac{m^{d / 2}}{\epsilon^{2}}\right) \equiv O\left(\frac{\sqrt{|D|}}{\epsilon^{2}}\right)$ samples and time.

- previous best for monotonicity testing: $\tilde{O}\left(\frac{m^{d-0.5}}{\epsilon^{4}}\right)$ [Bhattacharya-Fisher-Rubinfeld-Valiant'11]
- previous best for independence: $\mathrm{d}=2$, worst bounds [Batu et al.'01]

3. All bounds above are optimal

- i.e. matching lower bounds for both 1 and 2 via Le Cam Inequality.

4. Unified approach, computationally efficient tests, based on new $\chi^{2}$-tolerant tester
N.B. Contemporaneous work of [Canonne et al'2015] provide a different unified approach for testing structure but their results are suboptimal.

## Summary so far

- Hypothesis Testing in the small sample regime.


## Summary so far

- Hypothesis Testing in the small sample regime.



## Test

』
Pass/Fail?

## Summary so far

- Hypothesis Testing in the small sample regime.
- $p$ unknown distribution over some discrete set $D$
- $\mathcal{P}$ : set of distributions over $D$
$p$
- Given: $\epsilon, \boldsymbol{\delta}$, sample access to $p$
- Goal: w/ prob $\geq 1-\delta$ tell $p \in \mathcal{P}$ vs $\ell_{1}(p, \mathcal{P})>\epsilon$



## Test

』
Pass/Fail?

## Summary so far

- Hypothesis Testing in the small sample regime.
- punknown distribution over some discrete set $D$
- $\mathcal{P}$ : set of distributions over $D$
- Given: $\epsilon, \boldsymbol{\delta}$, sample access to $p$
- Goal: w/ prob $\geq 1-\delta$ tell $p \in \mathcal{P}$ vs $\ell_{1}(p, \mathcal{P})>\epsilon$
- Properties of interest: Is $p$ uniform? unimodal? logconcave? MHR? product measure?


## Test

【
Pass/Fail?

## Summary so far

- Hypothesis Testing in the small sample regime.
- p unknown distribution over some discrete set $D$
- $\mathcal{P}$ : set of distributions over $D$
$p$
- Given: $\epsilon, \boldsymbol{\delta}$, sample access to $p$
- Goal: w/ prob $\geq 1-\delta$ tell $p \in \mathcal{P}$ vs $\ell_{1}(p, \mathcal{P})>\epsilon$
- Properties of interest: Is $p$ uniform? unimodal? logconcave? MHR? product measure?
- All above properties can be tested $w / O\left(\frac{\sqrt{|D|}}{\epsilon^{2}} \cdot \log \frac{1}{\delta}\right)$ samples and time

Pass/Fail?

## Summary so far

- Hypothesis Testing in the small sample regime.
- p unknown distribution over some discrete set $D$
- $\mathcal{P}$ : set of distributions over $D$
- Given: $\epsilon, \boldsymbol{\delta}$, sample access to $p$
- Goal: w/ prob $\geq 1-\delta$ tell $p \in \mathcal{P}$ vs $\ell_{1}(p, \mathcal{P})>\epsilon$
- Properties of interest: Is $p$ uniform? unimodal? logconcave? MHR? product measure?
- All above properties can be tested $w / O\left(\frac{\sqrt{|D|}}{\epsilon^{2}} \cdot \log \frac{1}{\delta}\right)$ samples and time
- Unified approach based on modified Pearson's goodness of fit test: statistic $Z=\sum_{i \in D} \frac{\left(N_{i}-E_{i}\right)^{2}-N_{i}}{E_{i}}$


## Summary so far

- Hypothesis Testing in the small sample regime.
- p unknown distribution over some discrete set $D$
- $\mathcal{P}$ : set of distributions over $D$
- Given: $\epsilon, \boldsymbol{\delta}$, sample access to $p$
- Goal: w/ prob $\geq 1-\delta$ tell $p \in \mathcal{P}$ vs $\ell_{1}(p, \mathcal{P})>\epsilon$
- Properties of interest: Is $p$ uniform? unimodal? logconcave? MHR? product measure?
- All above properties can be tested $w / O\left(\frac{\sqrt{|D|}}{\epsilon^{2}} \cdot \log \frac{1}{\delta}\right)$ samples and time
- Unified approach based on modified Pearson's goodness
of fit test: statistic $Z=\sum_{i \in D} \frac{\left(N_{i}-E_{i}\right)^{2}-N_{i}}{E_{i}}$
- tight control for false positives: want to be able to both assert and reject the null hypothesis
- accommodate sublinear sample size


## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)


## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)



## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)
- Are bits/pixels independent?
- Our algorithms require $\Theta\left(\frac{2^{n / 2}}{\epsilon^{2}}\right)$ samples

- ...



## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)
- Are bits/pixels independent?
- Our algorithms require $\Theta\left(\frac{2^{n / 2}}{\epsilon^{2}}\right)$ samples
- Is source generating graphs over $n$ nodes Erdos-Renyi $G\left(n, \frac{1}{2}\right)$ ?
- Our algorithms require $\Theta\left(\frac{\binom{n}{2} / 2}{\epsilon^{2}}\right)$ samples


400 bit images


## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)
- Are bits/pixels independent?
- Our algorithms require $\Theta\left(\frac{2^{n / 2}}{\epsilon^{2}}\right)$ samples
- Is source generating graphs over $n$ nodes Erdos-Renyi $G\left(n, \frac{1}{2}\right)$ ?
- Our algorithms require $\Theta\left(\frac{2\binom{n}{2} / 2}{\epsilon^{2}}\right)$ samples

400 bit images


- Exponential dependence on $n$ unsettling, but necessary
- Lower bound exploits high possible correlation among bits


## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)


400 bit images

- Are bits/pixels independent?
- Our algorithms require $\Theta\left(\frac{2^{n / 2}}{\epsilon^{2}}\right)$ samples

- Is source generating graphs over $n$ nodes Erdos-Renyi $G\left(n, \frac{1}{2}\right)$ ?
- Our algorithms require $\Theta\left(\frac{2\binom{n}{2} / 2}{\epsilon^{2}}\right)$ samples
- Exponential dependence on $n$ unsettling, but necessary

- Lower bound exploits high possible correlation among bits
- Nature is not adversarial
- Often high dimensional systems have structure, e.g. Markov random fields fields (MRFs), graphical models (Bayes nets), etc


## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0,1\}^{n}$
- 0011010101 (sample 1)
- 0101001110 (sample 2)
- 0011110100 (sample 3)


400 bit images

- Are bits/pixels independent?
- Our algorithms require $\Theta\left(\frac{2^{n / 2}}{\epsilon^{2}}\right)$ samples

- Is source generating graphs over $n$ nodes Erdos-Renyi $G\left(n, \frac{1}{2}\right)$ ?
- Our algorithms require $\Theta\left(\frac{2\binom{n}{2} / 2}{\epsilon^{2}}\right)$ samples
- Exponential dependence on $n$ unsettling, but necessary

- Lower bound exploits high possible correlation among bits
- Nature is not adversarial
- Often high dimensional systems have structure, e.g. Markov random fields fields (MRFs), graphical models (Bayes nets), etc

Testing high-dimensional distributions with combinatorial structure?

## High-Dimensional Distn's

- Consider source generating $n$-bit strings $\in\{0.1\}^{n}$
- 00110101( [w/ Dikkala, Kamath'16]: If unknown $p$ is known to be an Ising
- 1010011. 
- 00111101 model, then poly $\left(n, \frac{1}{\epsilon}\right)$ samples suffice to test independence,
- ... goodness-of-fit. (extends to MRFs)
- Are bits/pixels independent?
- Our algorithms require $\Theta\left(\frac{2^{n / 2}}{\epsilon^{2}}\right)$ samples

- Is source generating graphs over $n$ nodes Erdos-Renyi $G\left(n, \frac{1}{2}\right)$ ?
- Our algorithms require $\Theta\left(\frac{2\binom{n}{2} / 2}{\epsilon^{2}}\right)$ samples
- Exponential dependence on $n$ unsettling, but necessary

- Lower bound exploits high possible correlation among bits
- Nature is not adversarial
- Often high dimensional systems have structure, e.g. Markov random fields fields (MRFs), graphical models (Bayes nets), etc

Testing high-dimensional distributions with combinatorial structure?

## Ising Model

- Statistical physics, computer vision, neuroscience, social science


## Ising Model

- Statistical physics, computer vision, neuroscience, social science
- Ising model:
- Probability distribution defined in terms of a graph $G=(V, E)$, edge potentials $\theta_{e}$, node potentials $\theta_{v}$



## Ising Model

- Statistical physics, computer vision, neuroscience, social science
- Ising model:
- Probability distribution defined in terms of a graph $G=(V, E)$, edge potentials $\theta_{e}$, node potentials $\theta_{v}$
- State space $\{ \pm 1\}^{V}$



## Ising Model

- Statistical physics, computer vision, neuroscience, social science
- Ising model:
- Probability distribution defined in terms of a graph $G=(V, E)$, edge potentials $\theta_{e}$, node potentials $\theta_{v}$
- State space $\{ \pm 1\}^{V}$

- High $\left|\theta_{e}\right|^{\prime} \mathrm{s} \Rightarrow$ strongly (anti-)correlated spins


## Ising Model: Strong vs weak ties


$\theta_{e}=1$

$\theta_{e}=0.5$

$\theta_{e}=0.25$

$\theta_{e}=0$

## Ising Model: Strong vs weak ties


$\theta_{e}=0.5$


## Ising Model: Strong vs weak ties

"low temperature regime"


$$
\theta_{e}=1 \quad \theta_{e}=0.5
$$



## Ising Model: Strong vs weak ties

## "low temperature regime"


"high temperature regime"

Phase Transition


Paramagnet

$\theta_{e}=1$

$\theta_{e}=0$

## Ising Model: Strong vs weak ties

"low temperature regime"


$$
\theta_{e}=1 \quad \theta_{e}=0.5
$$


"high temperature regime"


## $\theta_{v}=0$

## Ising Model: Strong vs weak ties

"low [w/ Dikkala, Kamath'16]: If unknown $p$ is known to be an Ising model, then poly $\left(n, \frac{1}{\epsilon}\right)$ samples suffice to test independence, goodness-of-fit.

- e.g. testing independence of ferromagnets (all $\theta_{e}>0$ ) needs $O\left(\frac{m}{\epsilon}\right)$ samples
- extends to MRFs


$$
\theta_{e}=1
$$


$\theta_{e}=0.5$

$\theta_{e}=0.25$

$\theta_{e}=0.125$

$\theta_{e}=0$

## e.g.4: Behavior in a Social Network



## e.g.4: Behavior in a Social Network



Question: Are adopted technologies a product distribution or are they far from being from a product distribution?

## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## The Menu

Motivation<br>Problem Formulation<br>Uniformity Testing, Goodness of Fit<br>Testing Properties of Distributions<br>Testing in High Dimensions<br>Conclusion

## Conclusions

- [w/ Acharya, Kamath'15]: Improved $\chi^{2}$-test, requiring $O\left(\frac{\sqrt{D}}{\epsilon^{2}}\right)$ samples
- implies testers of various distributional properties (independence, unimodality, logconcavity, etc) from same number of samples


## Conclusions

- [w/ Acharya, Kamath'15]: Improved $\chi^{2}$-test, requiring $O\left(\frac{\sqrt{D}}{\epsilon^{2}}\right)$ samples
- implies testers of various distributional properties (independence, unimodality, logconcavity, etc) from same number of samples
- Testing properties of high-dimensional distributions requires exponentially many samples


## Conclusions

- [w/ Acharya, Kamath'15]: Improved $\chi^{2}$-test, requiring $O\left(\frac{\sqrt{D}}{\epsilon^{2}}\right)$ samples
- implies testers of various distributional properties (independence, unimodality, logconcavity, etc) from same number of samples
- Testing properties of high-dimensional distributions requires exponentially many samples
- Making assumptions about the distribution being sampled gives leverage


## Conclusions

- [w/ Acharya, Kamath'15]: Improved $\chi^{2}$-test, requiring $O\left(\frac{\sqrt{D}}{\epsilon^{2}}\right)$ samples
- implies testers of various distributional properties (independence, unimodality, logconcavity, etc) from same number of samples
- Testing properties of high-dimensional distributions requires exponentially many samples
- Making assumptions about the distribution being sampled gives leverage
- [w/ Dikkala, Kamath'16]: Testing independence and goodness-of-fit in Ising models can be done with polynomially many samples


## Testing Combinatorial Structure



## Testing Combinatorial Structure

Is the phylogenic tree assumption true?


## Testing Combinatorial Structure

Is the phylogenic tree assumption true?

Sapiens-Neanderthal early interbreeding
[slatkin et al'13]


## Testing Combinatorial Structure

Is the phylogenic tree assumption true?

Sapiens-Neanderthal early interbreeding [slatkin et al'13]


## Testing from a Single Sample



## Testing from a Single Sample

- Given one social network, one brain, etc., how can we test the validity of a certain generative model?



## Testing from a Single Sample

- Given one social network, one brain, etc., how can we test the validity of a certain generative model?
- Get many samples from one sample?



## Testing from a Single Sample

- Given one social network, one brain, etc., how can we test the validity of a certain generative model?
- Get many samples from one sample?
- Ongoing with Rubinfeld



## Testing from a Single Sample

- Given one social network, one brain, etc., how can we test the validity of a certain generative model?
- Get many samples from one sample?
- Ongoing with Rubinfeld


> Thanks!

